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# Assessment of Antarctic Minke Whales using Statistical Catch-at-age Analysis 

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#### Abstract

Statistical Catch-at-age Analysis (SCAA) is applied to data for Antarctic minke whales. The SCAA model is spatially-structured, can model multiple stocks of minke whales, and can utilize several data types for parameter estimation. The application to Antarctic minke whales considers two stocks (I and P) in five areas which cover Antarctic Areas III-E to IV-W. The parameters of the model (annual deviations about the stock-recruitment relationship, density-dependence parameters (productivity and carrying capacity), and the parameters which determine growth by stock, age-specific natural mortality by stock, and vulnerability by area and 'fleet') are estimated by fitting the model to data on catches, catch-at-length, conditional age-at-length, and estimates of absolute and relative abundance. A reference case analysis is selected and sensitivity explored by varying the assumptions on which the reference case analysis is based. The reference case analysis is able to mimic all of the data sources adequately. Most of the analyses (reference and sensitivity) indicate that Antarctic minke whales in the assessed area increased from 1930 until the mid-1970s and have declined thereafter, with the extent of the decline greater for minke whales in Antarctic Areas III-E to V-W than for those further east. Natural mortality is consistently estimated to be higher for younger and older individuals. The estimates of MSYR ${ }_{1+}$ are $5.3 \%$ for minke whales in Antarctic Areas III-E to V-W and $3.6 \%$ for minke whales in Areas V-E and VI-W, but these estimates are less well determined than other model outputs, and quite sensitive to the assumptions on which the SCAA is based.


## KEYWORDS: CATCH-AT-AGE, ANTARCTIC MINKE WHALE, POPULATION MODEL; SOUTHERN HEMISPHERE

## INTRODUCTION

Two classes of stock assessment method have been proposed for application to Antarctic minke whales. One of these (ADAPT-VPA; Butterworth et al. [1996, 1999, 2002]) is based on the assumption that the catches-at-age at measured with limited error compared to the indices of abundance used to estimate the values for the parameters of the model. The other is Statistical Catch-at-Age Analysis (SCAA; Punt and Polacheck [2005, 2006, 2007, 2008]; Punt [2011]). In contrast to ADAPT-VPA, SCAA does not assume that the age-structure of the catches is measured with limited error, and can account for both sampling error and age-reading error ${ }^{1}$. In addition, the SCAA developed for Antarctic minke whales can account for multiple stocks in the assessed area, time-varying growth,

[^0]multiple areas, mixing of stocks in areas, environmental covariates, fleet-specific vulnerabilities, changes over time in the proportion of each stock in each area, and changes over time in vulnerability.

A question of considerable interest to the IWC Scientific Committee has been why the estimates of abundance from the IDCR/SOWER line transect survey exhibited declining trends in some Management Area. In addition, there has been interest in exploring the evidence for or against the hypothesis that there has been a recent decrease in carrying capacity for the Antarctic minke whales due to an increase in competition from other predators (IWC, 2005).

The paper applies SCAA to data for the Antarctic minke whales in Management Areas III-E, IV, V, and VI-W. It first outlines the mathematical specifications for the model and its associated estimation scheme. The paper then provides specifications for a 'reference' case analysis which uses all of the available index, catch length-composition, and conditional catch-at-age data. The reference case model considers five areas (III-E, IV, V-W, V-E, and VI-W), selected primarily because of the availability of data, and considers two stocks: the I stock - assumed to be found in areas III-E, IV, and V-W, and the P stock - assumed to be found in Areas V-E and VI-W. Full results for this reference case are provided based on suggestions for model outputs and fit diagnostics by the Scientific Committee. A series of sensitivity tests are outlined which examine the sensitivity of the results to the assumptions of the model, including that carrying capacity may have changed over time, and the weights assigned to each of the many data sources and penalties. Results for the sensitivity tests are restricted to a set of 'core' statistics to keep the volume of results to minimum. However, detailed results are available electronically for all of the sensitivity tests.

## MATHEMATICAL SPECIFICATIONS OF THE POPULATION MODEL

## A. Basic structure

The population dynamics model considers multiple stocks (indexed by ' $s$ ') and represents each stock using an age- and sex-structured model. 'Fleets' in this model are combinations of 'fleet' (Japan before 1987/88, Japan from 1987/88², and ex-Soviet Union) and area. Each 'fleet’ can have a different age- or length-specific vulnerability pattern (the combined effects of harvest selectivity and availability), which may change over time. Similarly growth, which depends on stock and sex, can change over time.

## B. The population dynamics model

Under the assumption that harvesting occurs instantaneously at the start of the year, the number of animals of stock $s$, $\operatorname{sex} g$ and age $a$ at the start of year $y, N_{y, a}^{g, s}$, is given by:

$$
N_{y, a}^{g, s}= \begin{cases}0.5 \tilde{N}_{y, 0}^{s} & \text { if } a=0  \tag{B.1}\\ \left(N_{y-1, a-1}^{g, s}-C_{y-1, a-1}^{g, s}\right) e^{-M_{a-1}^{s}} & \text { if } 1 \leq a \leq x-1 \\ \left(N_{y-1, x-1}^{g, s}-C_{y-1, x-1}^{g, s}\right) e^{-M_{x-1}^{s}}+\left(N_{y-1, x}^{g, s}-C_{y-1, x}^{g, s}\right) e^{-M_{x}^{s}} & \text { if } a=x\end{cases}
$$

[^1]where $\tilde{N}_{y, 0}^{s}$ is the number of births to stock $s$ at the start of year $y$ (the sex-ratio at birth is assumed to be 50:50), $C_{y, a}^{g, s}$ is the catch of animals of stock $s$, sex $g$ and age $a$ during year $y$, calculated as the sum of the catch of such animals over all fleets, i.e.:
\[

$$
\begin{equation*}
C_{y, a}^{g, s}=\sum_{f} C_{y, a}^{g, s, f} \tag{B.2}
\end{equation*}
$$

\]

$C_{y, a}^{g, s, f}$ is the catch of animals of stock $s$, sex $g$ and age $a$ by fleet $f$ during year $y, M_{a}^{s}$ is the instantaneous rate of natural mortality on animals of stock $s$ and age $a$ (assumed to be time-invariant), and $x$ is the plus-group (set equal to 54).

## C. Natural mortality-at-age

The relationship between natural mortality and age is taken to be piecewise linear:

$$
M_{a}^{s}= \begin{cases}\delta M^{s} & \text { if } a \leq a_{1}  \tag{C.1}\\ M^{s}\left[\delta+(1-\delta) \frac{\left(a-a_{1}\right)}{\left(a_{2}-a_{1}\right)}\right] & \text { if } a_{1}<a<a_{2} \\ M^{s} & \text { if } a_{2} \leq a \leq a_{3} \\ M^{s}\left[1+(\gamma-1) \frac{\left(a-a_{3}\right)}{\left(a_{4}-a_{3}\right)}\right] & \text { if } a_{3}<a<a_{4} \\ \gamma M^{s} & \text { if } a \geq a_{4}\end{cases}
$$

where $\delta M^{s}$ is the rate of natural mortality for animals of stock $s$ aged $a_{1}$ and younger, $M^{s}$ is the rate of natural mortality for animals of stock $s$ aged between $a_{2}$ and $a_{3}$, and $\gamma M^{s}$ is the rate of natural mortality for animals of stock $s$ aged $a_{4}$ and older.

## D. Births

The number of births to stock $s$ during year $y$ depends on the number of females that have reached the age-at-first-parturition at the start of year $y$ and the extent of densitydependence in pregnancy rate and infant survival ${ }^{3}$, i.e.:

$$
\begin{equation*}
B_{y}^{s}=B_{y}^{\mathrm{F}, s} f_{0}^{s} e^{A^{s}\left(1-B_{y}^{1+s /} / K_{y}^{1+s}\right)} e^{\varepsilon_{y}^{s}} \tag{D.1a}
\end{equation*}
$$

or

$$
\begin{equation*}
B_{y}^{s}=B_{y}^{\mathrm{F}, s} f_{0}^{s} e^{A^{s}\left(1-B_{y}^{1, s} s K_{y}^{1+s, s} e^{\varepsilon_{y}^{s}}\right.} \tag{D.1b}
\end{equation*}
$$

where $B_{y}^{\mathrm{F}, s}$ is the number of females of stock $s$ that have reached the age-at-firstparturition at the start of year $y$, i.e.:

[^2]\[

$$
\begin{equation*}
B_{y}^{\mathrm{F}, s}=\sum_{a=1}^{x} \beta_{y, a} N_{y, a}^{g, s} \tag{D.2}
\end{equation*}
$$

\]

$B_{y}^{1+s} \quad$ is the number of animals aged 1 and older in stock $s$ at the start of year $y$ :

$$
\begin{equation*}
B_{y}^{1+, s}=\sum_{g} \sum_{a=1}^{x} N_{y, a}^{g, s} \tag{D.3}
\end{equation*}
$$

$K_{y}^{1+, s}$ is the carrying capacity of stock $s$ (expressed in terms of the size of the $1^{+}$ component of the population) at the start of year $y, \beta_{y, a}$ is the proportion during year $y$ of animals of age $a$ that have reached the age-at-first-parturition, $f_{0}^{s}$ is the pregnancy rate / infant survival rate in absence of harvesting for stock $s, A^{s}$ is the resilience parameter for stock $s, \varepsilon_{y}^{s}$ is the logarithm of the ratio of the expected to actual number of births for stock $s$ during year $y$ (Equation D.1a only), and $\sigma_{R}$ is the standard deviation of $\varepsilon_{y}^{s}$.

Allowance is made for the possibility that carrying capacity has changed in a piecewise linear manner over the period considered in the analyses:

$$
K_{y}^{1+, s}= \begin{cases}K_{1930}^{1+, s} & \text { if } y \leq y_{1}  \tag{D.4}\\ K_{1930}^{1+, s}\left[1+\left(\tilde{K}_{I}^{1+, s}-1\right) \frac{\left(y-y_{1}\right)}{\left(y_{2}-y_{1}\right)}\right] & \text { if } y_{1}<y<y_{2} \\ K_{1930}^{1+, s}\left[\tilde{K}_{I}^{1+, s}+\left(\tilde{K}_{2002}^{1+, s}-\tilde{K}_{I}^{1+, s}\right) \frac{\left(y-y_{2}\right)}{\left(y_{3}-y_{2}\right)}\right] & \text { if } y_{2} \leq y<y_{3} \\ K_{1930}^{1+, s} \tilde{K}_{2002}^{1+, s} & \text { if } y \geq y_{3}\end{cases}
$$

where $K_{1930}^{1+, s}$ is the carrying capacity for stock $s$ from 1930 to year $y_{1}, \tilde{K}_{I}^{1+, s}$ is the ratio of the carrying capacity for stock $s$ in year $y_{2}$ to that in year $y_{1}$, and $\tilde{K}_{2002}^{1+s}$ is the ratio of the carrying capacity for stock $s$ in year $y_{3}$ to that in year $y_{1}$.

## E. Catches and vulnerability

The model-estimate of the catch of animals of stock $s$, sex $g$ and age $a$ by fleet $f$ during year $y$ depends on the number of animals of stock $s$, sex $g$ and age $a$, the exploitation rate by fleet $f$ on animals of sex $s$ during year $y$, the proportion of animals of stock $s$ in the area where fleet $f$ operates, and the relative vulnerability of animals of sex $g$ and age $a$ during year $y$ to fleet $f$ (assumed to be independent of stock). $C_{y, a}^{g, s, f}$ is computed using the formula:

$$
\begin{equation*}
C_{y, a}^{g, s, f}=\sum_{l} C_{y, a, l}^{g, s, f} \tag{E.1}
\end{equation*}
$$

where $C_{y, a, l}^{g, s, f}$ is the catch during year $y$ by fleet $f$ of animals of stock $s$, sex $g$ and age $a$ that are in length-class $l$ :

$$
C_{y, a, l}^{g, s, f}= \begin{cases}\tilde{S}_{a} S_{y, l}^{g, f} F_{y}^{g, f} X_{y, a, l}^{g, s} P_{y}^{A, s} N_{y, a}^{g, s} & \text { if vulnerability is length-specific }  \tag{E.2}\\ \tilde{S}_{a} S_{y, a}^{g, f} F_{y}^{g, f} X_{y, a, l}^{g, s} P_{y}^{A, s} N_{y, a}^{g, s} & \text { if vulnerability is age-specific }\end{cases}
$$

$S_{y, l}^{g, f}$ is the vulnerability of animals of sex $g$ and length $l$ to fleet $f$ during year $y, S_{y, a}^{g, f}$ is the vulnerability of animals of sex $g$ and age $a$ to fleet $f$ during year $y, \quad \tilde{S}_{a}$ is a factor to reduce the availability of animals of certain (younger) ages to the fishery, $F_{y}^{g, f}$ is the exploitation rate due to fleet $f$ on fully-selected (i.e. $S_{y, l}^{g, f} \rightarrow 1$; $S_{y, a}^{g, f} \rightarrow 1$ ) animals of sex $g$ during year $y, P_{y}^{A, s}$ is the proportion of stock $s$ that is in the area $A$ (where fleet $f$ is found in area $A$ ) during year $y$ (the model assumes that there is no sex- or age-structure to distribution),

$$
\begin{equation*}
P_{y}^{A, s}=\overline{P^{A, s}} e^{q_{y}^{A, s}} / \sum_{A^{\prime}} e^{q_{y}^{A_{y}^{\prime, s}}} \tag{E.3}
\end{equation*}
$$

$\overline{P^{A, s}}$ is the expected proportion of stock $s$ that is in area $A, \varphi_{y}^{A, s}$ is the deviation from the expected proportion for stock $s$ in area $A$ during year $y$, and $X_{y, a, l}^{q, s}$ is the proportion of animals of stock $s$, sex $g$ and age $a$ that are in length-class $l$ during year $y$.

Vulnerability by fleet is either assumed to be a function of length, fleet and sex, or a function of age, fleet and sex. The model has options which allow vulnerability to be uniform (Equations E.4a and E.5a), logistic (Equations E.4b and E.5b), or domed-shaped (Equations E.4c and E.5c), and can vary over time:

$$
\begin{align*}
& S_{y, l}^{g, f}=1  \tag{E.4a}\\
& S_{y, l}^{g, f}=\left(1+e^{-\ell \ln 19\left(L_{l}-L_{50, y}^{g, f}\right) / L_{\text {diff }}^{g, f}}\right)^{-1}  \tag{E.4b}\\
& S_{y, l}^{g, f}= \begin{cases}\exp \left(-\left(L_{l}-L_{50, y}^{g, f}\right)^{2} / L_{\text {left }}^{g, f}\right) & \text { if } L_{l} \leq L_{50, y}^{g, f} \\
\exp \left(-\left(L_{l}-L_{50, y}^{g, f}\right)^{2} / L_{\text {inght }}^{g, f}\right) & \text { otherwise }\end{cases}  \tag{E.4c}\\
& S_{y, a}^{g, f}=1  \tag{E.5a}\\
& S_{y, a}^{g, f}=\left(1+e^{-\ell \ln 19\left(a-a_{50, y}^{g, f}\right) / /_{\text {difif }}^{g, f}}\right)^{-1}  \tag{E.5b}\\
& S_{y, a}^{g, f}= \begin{cases}\exp \left(-\left(a-a_{50, y}^{g, f}\right)^{2} / a_{\text {left }}^{g, f}\right) & \text { if } a \leq a_{50, y}^{g, f} \\
\exp \left(-\left(a-a_{50, y}^{g, f}\right)^{2} / a_{\mathrm{right}}^{g, f}\right) & \text { otherwise }\end{cases} \tag{E.5c}
\end{align*}
$$

where $L_{50, y}^{g, f}$ is the length-at-50\%-vulnerability (logistic vulnerability) / length-at-fullvulnerability (dome-shaped vulnerability) for fleet $f$ during year $y$ for animals of $\operatorname{sex} g$ :

$$
\begin{equation*}
L_{50, y}^{g, f}=L_{50, y-1}^{g, f}+\delta_{y}^{g, f} \tag{E.6a}
\end{equation*}
$$

$a_{50, y}^{g, f}$ is the age-at-50\%-vulnerability (logistic vulnerability) / age-at-full-vulnerability (dome-shaped vulnerability) for fleet $f$ during year $y$ for animals of sex $g$ :

$$
\begin{equation*}
a_{50, y}^{g, f}=a_{50, y-1}^{g, f}+\delta_{y}^{g, f} \tag{E.6b}
\end{equation*}
$$

$\delta_{y}^{g, f}$ is the "vulnerability deviation" during year $y$ for fleet $f$ for animals of sex $g, L_{\text {diff }}^{g, f}$ is the width of the length-specific vulnerability ogive for fleet $f$ for animals of $\operatorname{sex} g, a_{\text {diff }}^{g, f}$ is the width of the age-specific vulnerability ogive for fleet $f$ for animals of $\operatorname{sex} g, L_{\text {left }}^{g, f}$ and $L_{\text {right }}^{g, f}$ are the parameters that determine the extent of dome-shapedness for the lengthspecific vulnerability ogive for fleet $f$ for animals of sex $g, a_{\text {left }}^{g, f}$ and $a_{\text {right }}^{g, f}$ are the parameters that determine the extent of dome-shapedness for the age-specific vulnerability ogive for fleet $f$ for animals of sex $g$, and $L_{l}$ is the length (in ft) corresponding to the mid-point of length-class $l$.

Time-dependence in vulnerability is modelled by allowing the length- (or age-)at-
 vulnerability ogive is the same each year, but the point at which vulnerability first equals 1 changes. Time-dependence in vulnerability was modelled in this way to avoid the overparameterization that might occur if allowance was also made for time-dependence in the parameters that determine the shape of the vulnerability ogive.

## F. Growth

The proportion of animals of stock $s$ and sex $g$ in age-class $a$ that are in length-class $l$ during $y, X_{y, a, l}^{g, s}$, is given by:
where $\Delta L$ is half of the width of each length-class $(0.5 \mathrm{ft}), \sigma_{\gamma}^{g, s}$ is the extent of variability about the growth curve for sex $g$ for animals of stock $s, \bar{L}_{y, a}^{g, s}$ is the expected length of an animal of stock $s$, sex $g$ and age $a$ during year $y$, assuming that length-at-age is governed by a von Bertalanffy growth curve and that the growth rate parameter $k_{y}^{g, s}$ varies over time, i.e.:

$$
\bar{L}_{y, a}^{g, s}= \begin{cases}L_{0}^{g, s} & \text { if } a=0  \tag{F.2}\\ L_{\infty}^{g, s}-\left(L_{\infty}^{g, s}-\bar{L}_{y-1, a-1}^{g, s}\right) e^{-k_{y-1}^{g, s}} & \text { otherwise }\end{cases}
$$

$L_{\infty}^{g, s}$ is the asymptotic length for animals of stock $s$ and $\operatorname{sex} g, k_{y}^{g, s}$ is the value of the Brody growth coefficient for animals of stock $s$ and sex $g$ during year $y$ :

$$
\begin{equation*}
k_{y}^{g, s}=k_{y-1}^{g, s} e^{v_{y}^{s}} \tag{F.3}
\end{equation*}
$$

$L_{0}^{g, s}$ is the length of an animal of age zero for animals of stock $s$ and sex $g$, and $v_{y}^{s}$ is the extent to which the growth rate changes from year $y$-1 to year $y$ for stock $s$.

## G. Initial conditions

The initial conditions ( $y_{1}=1930$ ) correspond to a population at its unexploited equilibrium level, i.e.:

$$
N_{y_{1}, a}^{g, s}= \begin{cases}0.5 \tilde{N}_{y, 0}^{s} & \text { if } a=0  \tag{G.1}\\ N_{y_{1}, a-1}^{g, s} e^{-M_{a-1}^{s}} & \text { if } 1 \leq a \leq x-1 \\ N_{y_{1}, x-1}^{g, s} e^{-M_{x-1}^{s}} /\left(1-e^{-M_{x}^{s}}\right) & \text { if } a=x\end{cases}
$$

where $\tilde{N}_{y, 0}^{s}$ is the expected number of calves in the absence of exploitation for stock $s$.
The value of the parameter $f_{0}^{s}$ is chosen so that the population remains in balance in the absence of exploitation, i.e.:

$$
\begin{equation*}
f_{0}^{s}=\left[\sum_{a=1}^{x-1} \beta_{y_{1}, a} e^{-\sum_{a=0}^{a-1} M_{a^{\prime}}^{s}}+\beta_{y_{1}, x^{-}} e^{-\sum_{a=0}^{x-1} M_{a^{\prime}}^{s}} /\left(1-e^{-M_{x}^{s}}\right)\right]^{-1} \tag{G.2}
\end{equation*}
$$

## OBJECTIVE FUNCTION

The objective function contains contributions from the data and from penalties on some of the parameters, i.e.:

$$
\begin{equation*}
L=\sum_{i} O_{i} \ell n L_{i}+\sum_{j} P_{j} \tag{H.1}
\end{equation*}
$$

where $\ell n L_{i}$ is the contribution of the $\mathrm{i}^{\text {th }}$ data source to the objective function, $P_{j}$ is the contribution of the $\mathrm{j}^{\text {th }}$ penalty term to the objective function, and $O_{i}$ is a factor to account for overdispersion.

The data included in the assessment are the annual catches (by fleet and sex), the estimates of abundance (IDCR and JARPA / JARPA II), the catch length-frequency data and the conditional age-at-length data, while there are penalties on the magnitudes of the deviations from the expected number of births (Equation D.1), on the inter-annual
deviations in the growth rate (Equation F.3), on the inter-annual variation in the proportion of the population in each area (see Equation E.3), and on the inter-annual deviations in vulnerability (Equation E.6). Each of these contributions is discussed in turn below. The equations listed below assume that data for each data-type are available for every year, and for all areas and fleets. This is not the case in reality, and the equations are modified appropriately in the absence of data for specific years, areas and fleets.

## H. 1 Catches

The contribution of the catches to the objective function is based on the assumption that any errors when measuring the catch are log-normally distributed ${ }^{4}$, i.e.:

$$
\begin{equation*}
\ell \mathrm{n} L_{1}=\sum_{y} \sum_{g} \sum_{f}\left\{\frac{1}{2 \sigma_{C}^{2}} \sum_{y}\left(\ell \mathrm{n} \tilde{C}_{y}^{g, f}-\ell \mathrm{n} C_{y}^{g, f}\right)^{2}\right\}+\text { Const } \tag{H.1}
\end{equation*}
$$

where $\tilde{C}_{y}^{g, f}$ is the actual catch by fleet $f$ of animals of sex $g$ during year $y, C_{y}^{g, f}$ is the model-estimate of total catch by fleet $f$ of animals of sex $g$ during year $y$ :

$$
\begin{equation*}
C_{y}^{g, f}=\sum_{s} \sum_{a} C_{y, a}^{g, s, f} \tag{H.2}
\end{equation*}
$$

$\sigma_{C}$ quantifies the extent of variation in catches.

## H. 2 Estimates of abundance

The estimates of abundance are assumed to be indices of $1^{+}$abundance, i.e.:

$$
\begin{equation*}
\ell \mathrm{n} L_{2}=\sum_{A} \sum_{y}\left\{\frac{1}{2\left(\tilde{\sigma}_{y}^{A}\right)^{2}}\left(\ell \operatorname{n} V_{y}^{A}-\ell \mathrm{n}\left(\chi^{A} B_{y}^{\text {Surv, } A}\right)\right)^{2}\right\}+\text { Const } \tag{H.3}
\end{equation*}
$$

where $V_{y}^{A}$ is the estimate of abundance for area $A$ and year $y, \chi^{A}$ is the bias factor for area $A, \tilde{\sigma}_{y}^{A}$ is the measurement error standard deviation, determined from the observation error standard deviation and the extent of additional variance, i.e.:

$$
\begin{equation*}
\left(\tilde{\sigma}_{y}^{A}\right)^{2}=\tau^{2}+\left(\phi_{y}^{A}\right)^{2} \tag{H.4}
\end{equation*}
$$

$\tau^{2}$ is the extent of additional variance, $\phi_{y}^{A}$ is the coefficient of variation of $V_{y}^{A}, B_{y}^{\text {Surv,A }}$ is the model-estimate of the total $(1+)$ abundance in area $A$ at the start of year $y$, i.e.:

$$
\begin{equation*}
B_{y}^{\text {Surv,A }}=\sum_{s} \sum_{g} \sum_{a>0} \sum_{l} P_{y}^{A, s} X_{y, a, l}^{g, s} S_{y, l}^{g, f^{*}} N_{y, a}^{g, s} \tag{H.5}
\end{equation*}
$$

$f^{*}$ is the fleet to which the abundance estimates pertain (set to the post-1987 Japanese fleet for the JARPA / JARPA II indices; set to uniform selectivity for the IDCR indices).

[^3]
## H. 3 Length-frequency data

The contribution of the length-frequency data to the objective function is based on the assumption that the catch by sex and fleet is taken multinomially from the vulnerable population, i.e.:

$$
\begin{equation*}
\ell \mathrm{nL} L_{3}=-\sum_{y} \sum_{f} \sum_{g} M_{y}^{g,} \sum_{l=l_{\min , y}}^{l_{\max , y}} \rho_{y, l}^{g, f} \ell \mathrm{n}\left(\hat{\rho}_{y, l}^{g, f} / \rho_{y, l}^{g, f}\right)+\text { Const } \tag{H.6}
\end{equation*}
$$

where $M_{y}^{g, f}$ is the effective sample size for the length-frequency data for animals of sex $g$ taken by fleet $f$ during year $y$ (set equal to the number of animals of sex $g$ taken by fleet $f$ during year $y$ for which information on length is available multiplied by an overdispersion factor), $\rho_{y, l}^{g, f}$ is the observed fraction of the catch of animals of sex $g$ taken by fleet $f$ during year $y$ that is in length-class $l, \hat{\rho}_{y, l}^{g, f}$ is the model-estimate of the fraction of the catch of animals of sex $g$ taken by fleet $f$ during year $y$ that is in lengthclass $l$ :

$$
\begin{equation*}
\hat{\rho}_{y, l}^{g, f}=\frac{\sum_{s} \sum_{a} C_{y, a, l}^{g, s, f}}{\sum_{s^{\prime}} \sum_{a^{\prime}} \sum_{l^{\prime}} C_{y, a^{\prime}, l^{\prime}}^{g, s^{\prime}}} \tag{H.7}
\end{equation*}
$$

Lengths $l_{\text {min, } y}$ and $l_{\text {max, } y}$ define the plus and minus groups for the length-frequency data for year $y$ (data and model-predictions for animals with length less than $l_{\text {min,y }}$ are pooled in the $l_{\text {min, } y}$ length-class while data and model-predictions for animals with length greater than $l_{\max , y}$ are pooled in the $l_{\max , y}$ length-class).

## H. 4 Conditional age-at-length data

The age data are included in the objective function under the assumption that sampling for age is multinomial conditioned on length, i.e.:

$$
\begin{equation*}
\ell \mathrm{n} L_{4}=\sum_{y} \sum_{f} \sum_{g} \sum_{l=l_{\min , y}}^{l_{\max , l}} \tilde{M}_{y, l}^{g, f} \sum_{a=a_{\min , y}}^{a_{\max , y}} \theta_{y, a, l}^{g, f} \ell n\left(\hat{\theta}_{y, a, l}^{g, f} / \theta_{y, a, l}^{g, f}\right)+\text { Const } \tag{H.8}
\end{equation*}
$$

where $\tilde{M}_{y, l}^{g, f}$ is the effective sample size for the age breakup of the animals of sex $g$ in length-class $l$ taken by fleet $f$ during year $y$ (set equal to the number of animals of sex $g$ in length-class $l$ taken by fleet $f$ during year $y$ for which information on length and age is available multiplied by an overdispersion factor), $\theta_{y, a, l}^{g, f}$ is the observed fraction of the catch of animals in length-class $l$ of sex $g$ taken by fleet $f$ during year $y$ that were aged to be age $a, \hat{\theta}_{y, a, l}^{g, f}$ is the model-estimate of the fraction of the catch of animals in length-class $l$ of $\operatorname{sex} g$ taken by fleet $f$ during year $y$ that were aged to be age $a$, i.e.:

$$
\begin{equation*}
\hat{\theta}_{y, a, l}^{g, f}=\frac{\sum_{s} \tilde{C}_{y, a, l}^{g, s, f}}{\sum_{s^{\prime}} \sum_{a^{\prime}} \tilde{C}_{y, a^{\prime}, l}^{g, s, f}} \tag{H.9}
\end{equation*}
$$

$\tilde{C}_{y, a, l}^{g, s, f}$ is the model-estimate of the number of animals of sex $g$ and stock $s$ in length-class $l$ caught by fleet $f$ during year $y$ that would have been aged to be age $a$ :

$$
\begin{equation*}
\tilde{C}_{y, a, l}^{g, s, f}=\sum_{a^{\prime}} Y_{a, a^{\prime}, y} C_{y, a^{\prime}, l}^{g, s, f} \tag{H.10}
\end{equation*}
$$

$Y_{a, a^{\prime}, y}$ is the fraction during year $y$ of animals of $\operatorname{sex} g$ and age $a$, that are aged to be age $a$ (the age-reading error matrix), i.e.:

$$
\begin{equation*}
Y_{a, a^{\prime}, y}=\int_{a-0.5}^{a+0.5} \frac{1}{\sqrt{2 \pi} \sigma_{a^{\prime}, y}^{\prime \prime}} e^{-\frac{\left(\lambda-\beta_{\left.a^{\prime}, y\right)^{2}}^{2\left(\sigma_{a}^{\prime}, y\right)^{2}}\right.}{}} d \lambda \tag{H.11}
\end{equation*}
$$

$\beta_{a, y}$ is the expected age based on age-readings for an animal of true age $a$ during year $y$, and $\sigma_{a, y}^{\prime \prime}$ is the standard error of the age-estimate for an animal of true age $a$ during year $y$.

Ages $a_{\text {min,y }}$ and $a_{\text {max,y }}$ define the plus and minus groups for the ageing data for year $y$, i.e. data and model-predictions for animals with age greater than $a_{\text {max, } y}$ are pooled at age $a_{\text {max }, y}{ }^{5}$ and those with age less than $a_{\text {min, } y}$ are pooled at age $a_{\text {min }, y}$.

## H. 5 Penalties

The penalty on the deviations from the expected number of births is based on the assumption that these deviations are log-normally distributed, i.e.:

$$
\begin{equation*}
P_{1}=\frac{1}{2 \sigma_{R}^{2}} \sum_{s} \sum_{y}\left(\varepsilon_{y}^{s}\right)^{2} \tag{H.12}
\end{equation*}
$$

The penalty on the changes over time in the vulnerability deviations is based on the assumption that these deviations are normally distributed, i.e.:

$$
\begin{equation*}
P_{2}=\frac{1}{2 \sigma_{s}^{2}} \sum_{g} \sum_{y} \sum_{f}\left(\delta_{y}^{g, f}\right)^{2} \tag{H.13}
\end{equation*}
$$

where $\sigma_{s}$ is the extent of inter-annual variation in the age-at- $50 \%$-vulnerability.
The penalty on the annual deviations in the proportion of each stock in each area is based on the assumption that these deviations are normally distributed, i.e.:

[^4]\[

$$
\begin{equation*}
P_{3}=\frac{1}{2 \sigma_{P}^{2}} \sum_{s} \sum_{y} \sum_{A}\left(\varphi_{y}^{A, S}\right)^{2} \tag{H.14}
\end{equation*}
$$

\]

where $\sigma_{P}$ is the extent of variation in the distribution of the stock.
The penalty on the inter-annual changes in the von Bertalanffy growth rate parameter is based on the assumption that these deviations are normally distributed, i.e.:

$$
\begin{equation*}
P_{4}=\frac{1}{2 \sigma_{k}^{2}} \sum_{s} \sum_{y}\left(v_{y}^{s}\right)^{2} \tag{H.15}
\end{equation*}
$$

## DATA UTILIZED

The data used when conducting assessments of the Antarctic minke whales consist of catches, abundance estimates, length frequency data, and conditional-at-age length data. The data include the catches and sighting survey information through the 2011/12 season.

## I. 1 Catches and length-frequency data

Table 1 lists the catches by sex, fleet (Japan and ex-Soviet Union) and Management Area (III-E, IV, V-W, V-E, and VI-W). The catches prior to 1971/72 are not allocated to fleet because these catches were taken by several nations. There is no information on the length-frequency of these catches so the vulnerability patterns for the years prior to 1971/72 are assumed to be equal to that for Japan in 1971/72, and the pre-1971/72 catches for Area V are split equally between Areas V-W and V-E. The results are unlikely to be sensitive to these assumptions given the small magnitude of the catches concerned.

## I. 2 Age-composition data

Age-composition data and hence conditional age-at-length data are only available for the Japanese catches. Table 2 lists the number of animals aged and the number of animals for which length data are available.

## I. 3 Indices of abundance

Table 3 lists the estimates of absolute abundance from the IDCR program (Okamura and Kitakado, 2012) and the indices of abundance based on the JARPA / JARPA II programme (Hakamada et al., in review; Hakamada, pers. comm).

## THE REFERENCE CASE ANALYSIS

The reference case analysis matches the no-USSR length data analysis of Punt (2011), i.e. the length-frequency data for the USSR fleet are ignored, and selectivity for the USSR and the Japanese fleet are assumed to be the same. The primary specifications of this analysis are (note that the letters before each section match those in the previous sections) as follows.

## C. Natural mortality

Natural mortality is assumed to change (in a piecewise linear fashion) at ages 3, 10, 30 and 35 (ages $a_{1}, a_{2}, a_{3}$ and $a_{4}$ in Equation C.1). Natural mortality for stock I is assumed to be a constant proportion of that for stock P.

## D. Births

1. Process error in births is modelled using Equation D.1a.
2. Carrying capacity is assumed to have changed in 1930, 1960 and 1980 (years $y_{1}$, $y_{2}$ and $y_{3}$ in Equation D.4).
3. The proportion of animals that have reached the age-at-first-parturition is defined by a logistic curve where $50 \%$ of animals reach first parturition at 8.5 years and $95 \%$ by 11.5 years. The first age at which an animal may reach first parturition is set equal to 3 years. These specifications were made for consistency with the analyses conducted by Butterworth and Punt (1999).
4. The variation in births, $\sigma_{R}$, is set to 0.2 .

## E. Catches and vulnerability

1. Vulnerability is length- rather than age-specific.
2. Vulnerability for the ex-USSR fleet is assumed to be the same as that for the Japanese fleet because of concerns regarding possible mis-reporting of lengths by the ex-USSR (IWC, 2011).
3. Vulnerability for the Japanese fleets (before 1987/88) is assumed to be timevarying double-normal function of length (See Equation E.4c).
4. Vulnerability for the JARPA / JARPA II fleet is assumed to be a logistic function of length and to the constant over space.
5. An age-specific availability factor is estimated for age 1 only.

## F. Growth

1. Values for the change in growth rate $\left(k_{y}^{g, s}\right)$ are estimated for each year from 1963/64 until 2011/12.

## G. Likelihood function

1. There is no survey bias for the IDCR/SOWER estimates (i.e. $\chi=1$ for the IDCR estimates), and no additional variance for either IDCR/SOWER or JARPA/JARPA II (i.e. $\tau^{2}=0$ in Equation H.3) ${ }^{6}$.
2. Separate survey bias parameters are estimated for the JAPRA / JARPA II indices in each of the five areas included in the analysis.
3. Age-reading error is modelled by assuming separate age-reading error matrices for each primary reader (Table 4)

$$
\begin{equation*}
\beta_{a}=\beta_{L}+\left(\beta_{H}-\beta_{L}\right) \frac{a}{70} ; \quad \sigma_{a}^{\prime \prime}=\sigma_{L}^{\prime \prime}+\left(\sigma_{H}^{\prime \prime}-\sigma_{L}^{\prime \prime}\right) \frac{a}{70} \tag{J.1}
\end{equation*}
$$

4. The minus- and plus-group ages when fitting to the conditional age-at-length data, $a_{\mathrm{min}, y}$ and $a_{\mathrm{max}, y}$, are set to 1 and 45 yr respectively.
5. The minus- and plus-group lengths, $l_{\min , y}$ and $l_{\max , y}$ for females are set to 22 ft and 32 ft for the period of commercial whaling, and 17 ft and 32ft for JARPA /
[^5]JARPA II while $l_{\text {min,y }}$ and $l_{\text {max }, y}$ for males are set to 22 ft and 31 ft for the period of commercial whaling and 17 ft and 31ft for JARPA / JARPA II. [These choices were made to avoid fitting the model to length-classes with few data.]
6. The ex-USSR length-frequency data are ignored because of possible misreporting.
7. The standard deviation of the logarithms of measuring the catch, $\sigma_{C}$, is set to 0.05 .
8. The parameter that determines the extent of variability in the vulnerability deviations, $\sigma_{s}$, is set to 10 .
9. The parameter that determines the extent of variability in the proportion of each stock in each area, $\sigma_{P}$, is set to 0.3.
10. The parameter that determines the extent of variability in growth rate, $\sigma_{k}$, is set to 0.1.
11. The values for the overdispersion parameters for length-frequency and conditional age-length data ( 0.70 and 0.85 respectively) were selected as outlined by Punt and Polacheck (2006).
The reference case choices for $\sigma_{C}, \sigma_{R}, \sigma_{S}, \sigma_{P}$, and $\sigma_{k}$ were made to force the model to replicate the catches closely, not to allow large deviations in births compared to those expected from the number of mature females, and to allow for large changes in vulnerability from one year to the next and in the proportion of the population in each area, if this suggested by the data.

Table 5 lists the estimable parameters of the reference case model and Table 6 lists the specifications for the sensitivity tests.

## DIAGNOSTIC STATISTICS AND SENSTIVITY TESTS

Table 7 outlines a list of diagnostic statistics and plots. The statistics and some of the plots were originally selected by the Scientific Committee, but subsequently modified based on requests for additional analyses. The plots summarize the key outputs from the assessments (e.g. Figures 1 and 2) as well as plots to understand those outputs somewhat better (e.g. Figures 3, 4, and 5).

## RESULTS

## Reference case analysis

The estimates of natural mortality indicate that natural mortality is highest for the youngest and oldest animals (Figure 2; Table 8). Natural mortality for the I stock is estimated to be higher at large age than for the $P$ stock ( $0.127 \mathrm{yr}^{-1}$ for animals 35 yr and older, compared to $0.106 \mathrm{yr}^{-1}$ ). Both stocks are estimated to have increased from 1930 until the early 1970s, with the I stock having declined subsequently thereafter (Figure 1). The increase in abundance is due primarily to an increase in recruitment owing to an increase in carrying capacity (Figure 2). Carrying capacity is estimated to have declined subsequently for both stocks, but the effect of this on recruitment and hence total population size is much smaller for the P than for the I stock (Figure 1). The I stock is estimated to have initially been larger than the P stock, but the P stock is currently the larger of the two stocks (Figure 1, Table 8). The recruitment deviations (Figure 1) suggest that there have been periods of good and poor recruitment since 1975 . MSYR $1_{1+}$ for the I
stock is estimated to be higher than that for the P stock (5.3\% compared to $3.6 \%$; Table 8)

The estimates of the proportions of the total population in each area changes over time, and vary inter-annually to better fit the abundance estimates, while there is also evidence for time-varying growth (Figure 3). The initial declining trends in the proportion of the total population in Areas III-E, IV and V-W is due to the faster estimated rate of increase for the P stock. The von Bertalanffy growth rate is estimated to have peaked in the mid-1980s and to have declined thereafter (Figure 3). The number of calves per-mature female remains below the biological maximum (Figure 4); previous versions of the SCAA led to unrealistically high values for the number of calves per mature female (see Punt [2011] for reasons for this change in result, which relate to the implementation of a penalty early during the minimization process).

Figures 5 and 6 illustrate how well the model is able to mimic the estimates of absolute and relative abundance given the estimated changes in abundance as well as inter-annual variation in the proportion of the population in each area. The confidence intervals for the abundance estimates generally intersect the population trajectory, indicating that the extent of process error in the proportion of the stocks in each area is sufficient to capture additional variance. Figure 7 provides plots which summarise the fits to the length-frequency data using three commonly-used summary plots (bubble plots of Pearson residuals, input versus inferred effective sample sizes, and fits to data aggregated over year). Figure 8 provides a more detailed summary for one of the fleet-sex combinations. A full set of length-frequency diagnostics is not included in this document to keep its length down (but is available on request from the first author). The model is able to mimic the length-frequency data on average (Figure 7, right panels), except perhaps for females in Area V-E. The assumed extent of overdispersion (0.70; dotted lines in Figure 7) is not ideal for any one fleet, but (as expected) the model captures the average relationship between effective and observed sample sizes. The fits in Figure 8 are amongst the best. As expected, the fits are best when sample sizes are high.

The final diagnostic plot (Figure 9) summarizes the fit to the conditional age-at-length data (again results are only shown for one fleet-sex combination). There are no major concerns with the fits to the mean ages-at-length, suggesting that the extent of timevarying growth is sufficient to mimic the changes in growth over time (see Figure 10 for the model-predicted growth curves).

Figures 11 and 12 show the length- and age-specific selectivity curves.

## Sensitivity tests

Assuming that process error impacts the density-dependent exponent (Equation D.1a; sensitivity test P1) leads qualitatively different trends for the P stock and a substantially larger size for the I stock (particularly between 1930 and the mid-1970s). However, the results for this sensitivity test should be considered with caution because it is unlikely that the minimization algorithm found the global minimum of the objective function. The marked difference in $\mathrm{MSYR}_{1+}$ for the two stocks, as well as the very high rate of natural mortality for sensitivity test P1 are surprising, but are perhaps a reflection of convergence to a local minimum (Table 7).

As expected, the rate of increase in abundance is less for sensitivity test P5 in which no allowance is made for carrying capacity to increase (Figure 13). The assumption of no change in carrying capacity leads to a much poor fit to the data (Table 9).

Sensitivity test P2, which does not allow for process error in spatial distribution, is supported (in an AIC sense ${ }^{7}$ ) over the reference case (Table 9), suggesting that including an additional variance parameter in the reference case analysis may provide an adequate way to address overdispersion. Sensitivity test P2, however, leads to essentially the same results as the reference case analysis (Figure 13), suggesting that changing the way overdisperson is handled will not impact the overall conclusions of this study. Allowing natural mortality to differ among stocks is supported by AIC (a $\Delta$ AIC is 6.2), but this difference in AIC is not enormous, especially given the large amount of data and the way those data are weighted.

The results are quite sensitive to changes to how selectivity is modelled. Allowing $g(0)$ to be less than 1 (sensitivity tests S5 and S6) does not change the quality of the fit of the model to the data (Table 9). However, as expected, population size is larger for these sensitivity tests (Figure 14). Sensitivity test S1 (selectivity depends on age not length) leads to qualitatively different results for the I stock. However, this sensitivity test led to a much poorer fit to the data ( $\Delta \mathrm{AIC}=1402$ ). Sensitivity test S 8 leads to a much poorer fit to the data (Table 8), although far fewer parameters are estimated. Overall, however, the estimated trends in abundance are similar for sensitivity tests S8 and the reference case analysis (Figure 14). Treating the JARPA indices of abundance as absolute (sensitivity test S9) leads to a poor fit to the data and overall to a lower trajectory of 1+ population size. This is not surprising because the estimates of the survey catchability coefficients (survey $q$ in Table 8) for the JARPA surveys are generally substantially smaller than 1 (Table 8).

The trends in abundance are insensitive to whether separate selectivity patterns are estimated for JARPA and JARPA II (sensitivity test S10; Figure 14). As expected, the fit of the model is better (lower negative log-likelihood) when allowance is made for selectivity for JARPA and JARPA II to differ (Table 9). However, sensitivity test S10 involves nine additional parameters (four selectivity parameters and five survey catchability coefficients) so an improved fit (lower negative log-likelihood) is expected.

The results are insensitive to changes to how the data are weighted (Table 8, Figure 15). The one exception is when ageing error is ignored. Historical abundance is much larger for this sensitivity test, and the decline in abundance during the 1970s is even more pronounced than is case for the reference case analysis. Table 9 does not list the values for the negative log-likelihood function for the "D" sensitivity tests because changing the weights assigned to the data changes the likelihood function, which makes the values for the negative log-likelihood non-comparable among sensitivity tests.

## DISCUSSION

Most of the analyses indicate that Antarctic minke whales in the assessed area increased from 1930 until the mid-1970s and have declined thereafter, with the extent of the decline greater for minke whales in Antarctic Areas III-E to V-W than those further east (Figures 13, 14 and 15). The trend in recent (last 20-odd) years is relatively flat or perhaps declining slowly. Natural mortality is consistently estimated to be higher for younger and older individuals. There are a few sensitivity tests for which these general conclusions do

[^6]not hold, but one of those sensitivity tests are cases in which there is evidence for nonconvergence ${ }^{8}$.

The reference case estimates of $\mathrm{MSYR}_{1+}$ are $5.3 \%$ for minke whales in Antarctic Areas III-E to V-W and $3.6 \%$ for minke whales in Areas V-E and VI-W, but these estimates are less well determined than other model outputs and quite sensitive to the assumptions on which the SCAA is based.

The estimates of the variances of the estimates of natural mortality are generally very low (CVs of $10 \%$ or less). These estimates should be interpreted with some caution given that they rely on asymptotic approximations and the weights assigned to the data / penalties.

The model on which the assessment is based is very complicated because it explicitly models five areas and two stocks, allows for time-varying selectivity, and consequently has over 1,000 parameters. The complexity arises because (a) some of the parameters (those related to natural mortality) are shared between stocks, necessitating that the two stocks are modelled simultaneously, and (b) the estimates of absolute abundance from IDCR are area-specific, which implies that the proportion of the population in each area needs to be modelled. In principle, the IDCR data could have been treated as relative indices of abundance, but then a constraint would have had to have been imposed on the survey catchability coefficients so that they sum to 1 over all areas in which a stock is found. The explicit allowance for spatial structure would also be needed if mixing of stocks was desired. If simulation testing of the assessment method is desired, it would be sensible to drop the area-structure and move to an parameterization where the catchability coefficients for IDCR add to 1 as this should substantial speed up the time it takes to fit the model.

The analyses are based on many assumptions, some of these have been explored in the tests of sensitivity. In general, the results are robust to those assumptions. Assumptions which could not be explored in detail in this study related to stock structure are perhaps of greatest concern. In particular, the analyses of this paper assume that there are two stocks of minke whales in Antarctic Areas III-E through VI-W and that there are no areas of mixing. Recently a transition area, where the two stocks mix with each other, has been suggested in part of Area VW (Pastene, 2006). The longitudinal sector of this transition area and the proportion of each stock in the transition area could change by year (Kitakado et al., 2012). The model has the capability to allow for such mixing but without agreed data to estimate mixing proportions, it is not currently possible to explore model configurations with mixing.

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[^7]facilitating this and the Institute of Cetacean Research and the Government of Japan for providing access to these data.

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Table 1
Catches by sex and Area
Table 1(a) - pre-1971/72 catches

| Year | Area III $^{*}$ |  | Area IV |  | Area V $^{*}$ |  | Area VI $^{*}$ |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | Female | Male | Female | Male | Female | Male | Female | Male |
| $1953 / 54$ | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 1 |
| $1954 / 55$ | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| $1955 / 56$ | 8 | 4 | 3 | 2 | 1 | 1 | 2 | 2 |
| $1956 / 57$ | 5 | 2 | 0 | 1 | 0 | 0 | 6 | 3 |
| $1957 / 58$ | 127 | 54 | 49 | 21 | 17 | 5 | 50 | 31 |
| $1958 / 59$ | 28 | 10 | 20 | 9 | 7 | 3 | 5 | 3 |
| $1959 / 60$ | 51 | 21 | 35 | 15 | 28 | 9 | 2 | 1 |
| $1960 / 61$ | 55 | 24 | 8 | 4 | 15 | 4 | 12 | 7 |
| $1961 / 62$ | 0 | 1 | 0 | 0 | 0 | 0 | 0 | 0 |
| $1962 / 63$ | 8 | 3 | 1 | 1 | 0 | 1 | 0 | 0 |
| $1963 / 64$ | 3 | 1 | 51 | 43 | 2 | 0 | 0 | 0 |
| $1964 / 65$ | 1 | 2 | 0 | 1 | 0 | 0 | 0 | 0 |
| $1965 / 66$ | 4 | 5 | 0 | 0 | 0 | 1 | 1 | 0 |
| $1966 / 67$ | 10 | 5 | 1 | 1 | 0 | 1 | 1 | 2 |
| $1967 / 68$ | 27 | 73 | 327 | 273 | 1 | 0 | 1 | 0 |
| $1968 / 69$ | 43 | 72 | 27 | 23 | 2 | 1 | 0 | 1 |
| $1969 / 70$ | 84 | 102 | 7 | 4 | 2 | 1 | 0 | 0 |
| $1970 / 71$ | 0 | 0 | 16 | 10 | 1 | 1 | 0 | 0 |
| * split equally between the $e a s t e r n ~ a n d ~ w e s t e r n ~ h a l f-A r e a s ~$ |  |  |  |  |  |  |  |  |

*     - split equally between the eastern and western half-Areas.

Table 1(b) - catches 1970/71 - 1986/87

| Year | Area III-E |  |  |  | Area IV |  |  |  | Area V-W |  |  |  | Area V-E |  |  |  | Area VI-W |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | Japan |  | Soviet Union |  | Japan |  | Soviet Union |  | Japan |  | Soviet Union |  | Japan |  | Soviet Union |  | Japan |  | Soviet Union |  |
|  | Female | Male | Female | Male | Female | Male | Female | Male | Female | Male | Female | Male | Female | Male | Female | Male | Female | Male | Female | Male |
| 1971/72 | 184 | 170 | 0 | 0 | 1728 | 929 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| 1972/73 | 0 | 0 | 351 | 298 | 975 | 1116 | 1172 | 1294 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| 1973/74 | 818 | 260 | 86 | 50 | 1282 | 761 | 1526 | 1000 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 3 | 10 | 0 | 0 |
| 1974/75 | 751 | 519 | 0 | 0 | 410 | 430 | 913 | 477 | 310 | 190 | 165 | 69 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| 1975/76 | 604 | 417 | 757 | 376 | 237 | 198 | 215 | 231 | 160 | 260 | 154 | 57 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| 1976/77 | 940 | 445 | 1176 | 313 | 432 | 518 | 251 | 399 | 495 | 515 | 375 | 82 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| 1977/78 | 614 | 398 | 656 | 133 | 353 | 128 | 359 | 123 | 316 | 298 | 189 | 27 | 22 | 32 | 0 | 0 | 83 | 156 | 74 | 110 |
| 1978/79 | 958 | 642 | 542 | 175 | 573 | 386 | 285 | 126 | 104 | 69 | 168 | 73 | 0 | 0 | 0 | 0 | 0 | 0 | 5 | 5 |
| 1979/80 | 395 | 308 | 641 | 132 | 482 | 1048 | 202 | 129 | 113 | 383 | 687 | 161 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| 1980/81 | 292 | 327 | 343 | 275 | 664 | 529 | 841 | 352 | 330 | 105 | 132 | 114 | 156 | 34 | 335 | 42 | 99 | 100 | 10 | 48 |
| 1981/82 | 71 | 188 | 485 | 380 | 1043 | 582 | 0 | 0 | 779 | 369 | 0 | 0 | 11 | 18 | 0 | 0 | 67 | 218 | 0 | 0 |
| 1982/83 | 0 | 0 | 638 | 464 | 490 | 530 | 741 | 207 | 1480 | 416 | 0 | 0 | 0 | 0 | 0 | 0 | 170 | 137 | 319 | 150 |
| 1983/84 | 0 | 0 | 105 | 158 | 518 | 589 | 631 | 357 | 945 | 436 | 0 | 0 | 56 | 8 | 0 | 0 | 349 | 126 | 0 | 0 |
| 1984/85 | 0 | 0 | 377 | 142 | 364 | 137 | 659 | 328 | 573 | 337 | 0 | 0 | 0 | 0 | 0 | 0 | 92 | 277 | 0 | 0 |
| 1985/86 | 0 | 0 | 0 | 0 | 292 | 222 | 664 | 229 | 670 | 343 | 0 | 0 | 0 | 0 | 0 | 0 | 97 | 300 | 0 | 0 |
| 1986/87 | 0 | 0 | 41 | 21 | 321 | 193 | 628 | 322 | 851 | 162 | 0 | 0 | 0 | 0 | 0 | 0 | 285 | 129 | 0 | 0 |

Table 1(c) - catches by Japan post 1986/87

| Year | Area III-E |  | Area IV |  | Area V-W |  | Area V-E |  | Area VI-W |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | Female | Male | Female | Male | Female | Male | Female | Male | Female | Male |
| 1987/88 | 0 | 0 | 119 | 153 | 0 | 0 | 0 | 0 | 0 | 0 |
| 1988/89 | 0 | 0 | 0 | 0 | 0 | 0 | 151 | 85 | 0 | 0 |
| 1989/90 | 0 | 0 | 142 | 184 | 0 | 0 | 0 | 0 | 0 | 0 |
| 1990/91 | 0 | 0 | 0 | 0 | 77 | 110 | 68 | 54 | 14 | 0 |
| 1991/92 | 0 | 0 | 123 | 165 | 0 | 0 | 0 | 0 | 0 | 0 |
| 1992/93 | 0 | 0 | 0 | 0 | 87 | 118 | 53 | 45 | 20 | 4 |
| 1993/94 | 0 | 0 | 130 | 200 | 0 | 0 | 0 | 0 | 0 | 0 |
| 1994/95 | 0 | 0 | 0 | 0 | 27 | 113 | 103 | 87 | 0 | 0 |
| 1995/96 | 41 | 68 | 126 | 204 | 0 | 1 | 0 | 0 | 0 | 0 |
| 1996/97 | 0 | 0 | 0 | 0 | 72 | 55 | 80 | 77 | 82 | 74 |
| 1997/98 | 36 | 75 | 123 | 204 | 0 | 0 | 0 | 0 | 0 | 0 |
| 1998/99 | 0 | 0 | 0 | 0 | 88 | 95 | 34 | 111 | 20 | 41 |
| 1999/00 | 46 | 63 | 160 | 170 | 0 | 0 | 0 | 0 | 0 | 0 |
| 2000/01 | 0 | 0 | 0 | 0 | 45 | 95 | 73 | 87 | 64 | 76 |
| 2001/02 | 56 | 54 | 183 | 147 | 0 | 0 | 0 | 0 | 0 | 0 |
| 2002/03 | 0 | 0 | 0 | 0 | 46 | 54 | 116 | 114 | 43 | 67 |
| 2003/04 | 48 | 62 | 192 | 138 | 0 | 0 | 0 | 0 | 0 | 0 |
| 2004/05 | 0 | 0 | 0 | 0 | 47 | 35 | 137 | 75 | 79 | 67 |
| 2005/06 | 66 | 64 | 264 | 309 | 59 | 89 | 2 | 0 | 0 | 0 |
| 2006/07 | 0 | 0 | 0 | 0 | 0 | 0 | 233 | 88 | 118 | 66 |
| 2007/08 | 79 | 150 | 141 | 94 | 58 | 29 | 0 | 0 | 0 | 0 |
| 2008/09 | 0 | 0 | 0 | 0 | 57 | 77 | 106 | 67 | 141 | 231 |
| 2009/10 | 143 | 102 | 63 | 40 | 58 | 90 | 5 | 5 | 0 | 0 |
| 2010/11 | 0 | 0 | 0 | 0 | 0 | 0 | 20 | 26 | 88 | 36 |
| 2011/12 | 0 | 0 | 0 | 0 | 50 | 49 | 16 | 19 | 101 | 31 |

Table 2
Summary of the compositional data (number of animals aged and number of animals measured from the catch by Japan)

| Year | Area III-E |  |  |  | Area IV |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | Age-composition |  | Length-frequency <br> (Japan) |  | Age-composition |  | Length-frequency <br> (Japan) |  |
|  | F | M | F | M | F | M | F | M |
| 1971/72 | 12 | 6 | 184 | 170 | 487 | 235 | 1728 | 929 |
| 1972/73 | 0 | 0 | 0 | 0 | 413 | 418 | 975 | 1116 |
| 1973/74 | 250 | 85 | 818 | 260 | 436 | 272 | 1282 | 761 |
| 1974/75 | 468 | 285 | 751 | 519 | 235 | 257 | 410 | 430 |
| 1975/76 | 169 | 100 | 604 | 417 | 114 | 71 | 237 | 198 |
| 1976/77 | 352 | 146 | 940 | 445 | 156 | 168 | 432 | 518 |
| 1977/78 | 254 | 148 | 614 | 398 | 194 | 67 | 353 | 128 |
| 1978/79 | 643 | 439 | 958 | 642 | 428 | 274 | 573 | 386 |
| 1979/80 | 283 | 211 | 395 | 308 | 355 | 781 | 482 | 1048 |
| 1980/81 | 252 | 250 | 292 | 327 | 544 | 417 | 664 | 529 |
| 1981/82 | 62 | 149 | 71 | 188 | 864 | 491 | 1043 | 582 |
| 1982/83 | 0 | 0 | 0 | 0 | 392 | 398 | 490 | 530 |
| 1983/84 | 0 | 0 | 0 | 0 | 380 | 385 | 518 | 589 |
| 1984/85 | 0 | 0 | 0 | 0 | 303 | 110 | 364 | 137 |
| 1985/86 | 0 | 0 | 0 | 0 | 247 | 188 | 292 | 222 |
| 1986/87 | 0 | 0 | 0 | 0 | 293 | 177 | 321 | 193 |
| 1987/88 | 0 | 0 | 0 | 0 | 99 | 135 | 119 | 153 |
| 1988/89 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| 1989/90 | 0 | 0 | 0 | 0 | 118 | 155 | 142 | 184 |
| 1990/91 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| 1991/92 | 0 | 0 | 0 | 0 | 102 | 143 | 123 | 165 |
| 1992/93 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| 1993/94 | 0 | 0 | 0 | 0 | 102 | 173 | 130 | 200 |
| 1994/95 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| 1995/96 | 36 | 54 | 41 | 67 | 98 | 176 | 126 | 204 |
| 1996/97 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| 1997/98 | 36 | 63 | 36 | 75 | 91 | 168 | 123 | 204 |
| 1998/99 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| 1999/00 | 34 | 48 | 46 | 63 | 145 | 147 | 160 | 170 |
| 2000/01 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| 2001/02 | 45 | 49 | 56 | 54 | 157 | 131 | 183 | 147 |
| 2002/03 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| 2003/04 | 35 | 53 | 48 | 62 | 169 | 111 | 192 | 138 |
| 2004/05 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| 2005/06 | 45 | 43 | 66 | 64 | 204 | 248 | 264 | 309 |
| 2006/07 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| 2007/08 | 64 | 115 | 79 | 150 | 118 | 66 | 141 | 94 |
| 2008/09 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| 2009/10 | 120 | 79 | 143 | 102 | 56 | 27 | 63 | 40 |
| 2010/11 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| 2011.12 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |

(Table 2 Continued)

| Year | Area V-W |  |  |  | Area V-E |  |  |  | Area VI-W |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | Age-composition |  | Length-frequency (Japan) |  | Age-composition |  | Length-frequency (Japan) |  | Age-composition |  | Length-frequency <br> (Japan) |  |
|  | F | M | F | M | F | M | F | M | F | M | F | M |
| 1971/72 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| 1972/73 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| 1973/74 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 1 | 4 | 3 | 10 |
| 1974/75 | 145 | 54 | 310 | 190 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| 1975/76 | 66 | 132 | 160 | 260 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| 1976/77 | 263 | 237 | 495 | 515 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| 1977/78 | 209 | 191 | 316 | 298 | 19 | 24 | 22 | 32 | 23 | 45 | 83 | 156 |
| 1978/79 | 93 | 54 | 104 | 69 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| 1979/80 | 81 | 257 | 113 | 383 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| 1980/81 | 257 | 71 | 330 | 105 | 119 | 28 | 156 | 34 | 68 | 78 | 99 | 100 |
| 1981/82 | 548 | 256 | 779 | 369 | 10 | 15 | 11 | 18 | 49 | 157 | 67 | 218 |
| 1982/83 | 1109 | 303 | 1480 | 416 | 0 | 0 | 0 | 0 | 130 | 98 | 170 | 137 |
| 1983/84 | 717 | 316 | 945 | 436 | 48 | 6 | 56 | 8 | 279 | 87 | 349 | 126 |
| 1984/85 | 485 | 274 | 573 | 337 | 0 | 0 | 0 | 0 | 79 | 240 | 92 | 277 |
| 1985/86 | 596 | 311 | 670 | 343 | 0 | 0 | 0 | 0 | 77 | 250 | 97 | 300 |
| 1986/87 | 743 | 143 | 851 | 162 | 0 | 0 | 0 | 0 | 242 | 112 | 285 | 129 |

(Table 2 Continued)

| Year | Area V-W |  |  |  | Area V-E |  |  |  | Area VI-W |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | Age-composition |  | Length-frequency <br> (Japan) |  | Age-composition |  | Length-frequency (Japan) |  | Age-composition |  | Length-frequency <br> (Japan) |  |
|  | F | M | F | M | F | M | F | M | F | M | F | M |
| 1987/88 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| 1988/89 | 0 | 0 | 0 | 0 | 122 | 64 | 151 | 85 | 0 | 0 | 0 | 0 |
| 1989/90 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| 1990/91 | 67 | 101 | 77 | 110 | 65 | 50 | 68 | 54 | 14 | 0 | 14 | 0 |
| 1991/92 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| 1992/93 | 78 | 112 | 87 | 118 | 49 | 42 | 53 | 45 | 17 | 4 | 20 | 4 |
| 1993/94 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| 1994/95 | 25 | 99 | 27 | 113 | 88 | 72 | 103 | 87 | 0 | 0 | 0 | 0 |
| 1995/96 | 0 | 1 | 0 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| 1996/97 | 64 | 51 | 72 | 55 | 69 | 70 | 80 | 77 | 72 | 66 | 82 | 74 |
| 1997/98 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| 1998/99 | 72 | 80 | 88 | 95 | 32 | 90 | 34 | 111 | 12 | 38 | 20 | 41 |
| 1999/00 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| 2000/01 | 34 | 81 | 45 | 95 | 62 | 78 | 73 | 87 | 59 | 62 | 64 | 76 |
| 2001/02 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| 2002/03 | 37 | 48 | 46 | 54 | 100 | 98 | 116 | 114 | 32 | 54 | 43 | 67 |
| 2003/04 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| 2004/05 | 43 | 35 | 47 | 35 | 120 | 65 | 137 | 75 | 67 | 58 | 79 | 67 |
| 2005/06 | 45 | 70 | 59 | 89 | 2 | 0 | 2 | 0 | 0 | 0 | 0 | 0 |
| 2006/07 | 0 | 0 | 0 | 0 | 214 | 78 | 233 | 88 | 106 | 61 | 118 | 66 |
| 2007/08 | 54 | 21 | 58 | 29 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| 2008/09 | 45 | 64 | 57 | 77 | 95 | 55 | 106 | 67 | 111 | 198 | 141 | 231 |
| 2009/10 | 41 | 71 | 58 | 90 | 2 | 4 | 5 | 5 | 0 | 0 | 0 | 0 |
| 2010/11 | 0 | 0 | 0 | 0 | 16 | 22 | 20 | 26 | 75 | 34 | 88 | 36 |
| 2011/12 | 38 | 44 | 50 | 49 | 13 | 17 | 16 | 19 | 96 | 29 | 101 | 31 |

Table 3
The estimates of abundance
(a) IDCR Estimates

| Year | Estimate | Year | Estimate |
| :---: | :---: | :---: | :---: |
| Area III-E |  | Area IV |  |
| 1987/88 | $11782(0.440)$ | $1988 / 89$ | $46763(0.169)$ |
| 1994/95 | $34659(0.237)$ | $1998 / 99$ | $55873(0.341)$ |
| Area V-W |  | Area V-E |  |
| 1985/86 | $105951(0.159)$ | $1985 / 86$ | $154658(0.189)$ |
| 2001/02 | $43640(0.139)$ | $2003 / 04$ | $136457(0.134)$ |
| Area VI-W |  |  |  |
| 1990/91 | $20438(0.271)$ |  |  |
| 1995/96 | $48206(0.177)$ |  |  |

(b) JARPA / JARPA II indices of relative abundance

| Area III-E |  | Area IV |  | Area V-W |  |
| :---: | :---: | :---: | :---: | :---: | :---: |
| Year | Estimate | Year | Estimate | Year | Estimate |
| 1995/96 | 9614 (0.220) | 1989/90 | 29993 (0.228) | 1990/91 | 35108 (0.220) |
| 1997/98 | 5566 (0.367) | 1991/92 | 32418 (0.396) | 1992/93 | 22404 (0.350) |
| 1999/00 | 12404 (0.615) | 1993/94 | 27780 (0.147) | 1994/95 | 12805 (0.275) |
| 2001/02 | 44801 (0.582) | 1995/96 | 31601 (0.198) | 1996/97 | 15540 (0.293) |
| 2003/04 | 18927 (0.355) | 1997/98* | 16590 (0.277) | 1998/99 | 56927 (0.538) |
| 2005/06 | $292610.379)$ | 1999/00 | 43673 (0.125) | 2000/01 | 19603 (1.031) |
| 2007/08 | 9406 (0.291) | 2001/02 | 30269 (0.218) | 2002/03 | 66544 (0.234) |
|  |  | 2003/04 | 34701 (0.373) | 2004/05 | 27544 (0.243) |
|  |  | 2005/06 | 62979 (0.334) | 2005/06 | 45541 (0.701) |
|  |  | 2007/08 | 14739 (0.570) | 2006/07 | 30422 (0.534) |
|  |  |  |  | 2007/08 | 45157 (0.408) |
|  |  |  |  | 2008/09 | 10534 (0.638) |
|  | , V-E |  | VI-W |  |  |
| Year | Estimate | Year | Estimate |  |  |
| 1990/91 | 65638 (0.353) | 1996/97 | 7530 (0.258) |  |  |
| 1992/93 | 43743 (0.376) | 1998/99 | 20166 (0.280) |  |  |
| 1994/95 | 100771 (0.254) | 2000/01 | 12571 (0.257) |  |  |
| 1996/97 | 82155 (0.300) | 2002/03 | 10543 (0.218) |  |  |
| 1998/99 | 43037 (0.286) | 2004/05 | 14843 (0.228) |  |  |
| 2000/01 | 88274 (0.255) | 2006/07 | 17206 (0.556) |  |  |
| 2002/03 | 59961 (0.244) | 2008/09 | 26364 (0.226) |  |  |
| 2004/05 | 45177 (0.229) |  |  |  |  |
| 2006/07 | 26418 (0.188) |  |  |  |  |
| 2008/09 | 97563 (0.193) |  |  |  |  |

* Survey covered only a small part of Prydz Bay

Table 4
Parameters which determine the age-reading error matrix. These values correspond to the "Lockyer unbiased" analysis of Kitakado and Punt (2010) \{model 3\}. The values for $L$ and $H$ are 0 and 70 respectively.

| Year | Reader | $\beta_{L}$ | $\beta_{H}$ | $\sigma_{L}^{\prime \prime}$ | $\sigma_{H}^{\prime \prime}$ |
| :--- | :---: | :---: | :---: | :---: | :---: |
| $1971 / 72-1979 / 80$ | Masaki | 3.0837 | 58.7940 | 1.5531 | 1.5531 |
| $1980 / 81-1989 / 90$ | Kato | 2.4545 | 56.0060 | 0.5391 | 7.3718 |
| $1990 / 91-1991 / 92$ | Zenitani | 1.0300 | 62.8530 | 0.4637 | 3.6614 |
| $1992 / 93$ | Kato | 2.4545 | 56.0060 | 0.5391 | 7.3718 |
| $1993 / 94-2004 / 05$ | Zenitani | 1.0300 | 62.8530 | 0.4637 | 3.6614 |
| $2005 / 06-2011 / 12$ | Bando | 1.6355 | 63.6440 | 0.6588 | 3.4128 |

## Table 5

The estimable parameters of the population dynamics model and the objective function.

| Parameter | Number of parameters |  |  |
| :--- | :---: | :---: | :---: |
|  | Stock I | Stock P |  |
| Calves in the absence of exploitation, $B_{0}$ | 1 | 1 |  |
| Natural mortality: $M_{0}, M_{1} / M_{0}, M_{2} / M_{1}$ |  | 4 |  |
| Resilience, $A$ |  |  |  |
| Survival deviation, $\varepsilon_{y}$ | 1 | 1 |  |
| Expected proportion in each Area, $\overline{P^{s, A}}$ | 83 | 83 |  |
| Annual deviations about the expected proportions in each area, $\varphi_{y}^{A}$ | 2 | 1 |  |
| Exploitation rate by year, sex and fleet, $F_{y}^{g, f}$ | 69 | 26 |  |
| Changes in carrying capacity, $\tilde{K}_{I}^{1+}, \tilde{K}_{2002}^{1+}$ | 291 | 125 |  |
| Parameters of the growth curve, $L_{\infty}^{g}, k^{g}, t_{0}^{g}, \sigma_{\gamma}^{g}$ | 2 | 2 |  |
| Inter-annual deviations in growth rate, $v_{y}$ | 8 | 8 |  |
| Parameters to define vulnerability, $L_{50, y}^{g, f}, \delta_{y}^{g, f}, L_{d f f}^{g, f}, L_{\text {left }}^{g, f}$, | $14+72$ | 98 |  |
| $L_{\text {right }}^{g, f} \tilde{S}_{a}$ |  | $8+22$ |  |
| JARPA survey bias, $\chi$ |  |  |  |
| Total |  |  |  |

## Table 6

The sensitivity tests.

| Run | Description |
| :---: | :---: |
| 0 | Reference case |
| Related to the population dynamics model |  |
| P1 | Equation D.1b instead of Equation D.1a |
| P2 | The proportion of each stock in each area is time-invariant |
| P3 | No time-varying growth |
| P4 | Natural mortality is independent of stock |
| P5 | Carrying capacity ( $K$ ) is time-invariant |
| P6 | Carrying capacity changes in an auto-regressive manner |
| P7 | Carrying capacity changes in 1960, 1980, and 2002 |
| P8 | JARPA indices pertain to $1+$ abundance |
| Related to selectivity patterns |  |
| S1 | Selectivity is age- rather than length-specific |
| S2 | The length frequency and age composition data for the years until1973/74 are down-weighted by $90 \%$ |
| S3 | JARPA / JARPA II vulnerability is constant |
| S4 | JARPA / JARPA II vulnerability is constant (and time-invariant); carrying capacity is time-invariant |
| S5 | IDCR/SOWER estimates are assumed to be negatively biased, $\mathrm{g}(0)=0.60$ |
| S6 | IDCR/SOWER estimates are assumed to be negatively biased, $\mathrm{g}(0)=0.80$ |
| S7 | Ignore JARPA / JARPA II abundance estimates |
| S8 | Time-invariant fishery selectivity |
| S9 | JARPA / JARPA II abundance estimates are absolute |
| S10 | Separate selectivity patterns for JARPA / JARPA II |
| Related to data selection and data weighting |  |
| D1 | Ignore JARPA II data (the area covered by JARPA II differs from that for JARPA) |
| D2 | Ignore ageing error |
| D3 | Decrease $\sigma_{s}$ by $50 \%$ |
| D4 | Increase $\sigma_{s}$ by $50 \%$ |
| D5 | Decrease $\sigma_{\mathrm{r}}$ to 0.10 |
| D6 | Increase $\sigma_{\mathrm{r}}$ to 0.30 |
| D7 | Decrease $\sigma_{p}$ by $50 \%$ |
| D8 | Increase $\sigma_{\mathrm{p}}$ by $50 \%$ |
| D9 | Decrease $\sigma_{\mathrm{k}}$ by $50 \%$ |
| D10 | Increase $\sigma_{\mathrm{k}}$ by $50 \%$ |
| D11 | Double weight on length data |
| D12 | Halve weight on length data |
| D13 | Double weight on conditional age-at-length data |
| D14 | Halve weight on conditional age-at-length data |

## Table 7

The summary statistics and plots

```
(a) Statistics
    brec,1945-68 - slope of the linear regression of the estimates of the logarithms of the numbers of recruits (age 1
    animals) on time (1945-68).
    brec,1968-88 - slope of the linear regression of the estimates of the logarithms of the numbers of recruits on
    time (1968-88).
    brec,1988-End
    time (1988-last year).
    N Not,1945-68 - slope of the linear regression of the estimates of the logarithms of the numbers of 1+ animals on
    time (1945-68).
    N Not,1968-88 - slope of the linear regression of the estimates of the logarithms of the numbers of 1+ animals on
    time (1968-88).
    N Not,1988-Endr - slope of the linear regression of the estimates of the logarithms of the numbers of 1+ animals
    on time (1988-last year).
    N Nnd-5,1
    K}1930\mathrm{ - Carrying capacity in 1930.
    K}\mp@subsup{K}{200}{}/\mp@subsup{K}{1960}{}\mathrm{ - ratio of K in 2000 to that in 1960.
    K}\mp@subsup{K}{1960}{}/\mp@subsup{K}{1930}{*}\mathrm{ - ratio of K in 1960 to that in 1930.
    Natural mortality (ages 0-3, 10-30, 35+)
    Average proportions in each management area
    Survey q for JARPA.
MSYR (1+)
```


## (b) Plots

## Assessment outputs

Total (1+) population size versus year (by stock and by area)
Age 1 animals (recruits) versus time
Carrying capacity versus year
Natural mortality versus age
Number of females beyond the age-at-first parturition
Number of calves as a fraction of the number of females beyond the age-at-first parturition
Selectivity-at-age
Selectivity-at-length
Brody growth coefficient versus year

## Diagnostic plots

Survey estimates of abundance from IDCR with the associated model predictions (by area)
Survey estimates of abundance from JARPA with the associated model predictions (by area)
Observed and model-predicted catches (by fleet and sex)
Observed and model-predicted length-frequencies (by fleet, sex and year)
Input and estimated effective sample sizes for the length-frequency data (by fleet and sex)
Fits to the length-frequency data in the form of bubble plots (by fleet and sex)
Fits to the summed length-frequency data (by fleet and sex)

## Table 8

Results of the reference case analysis and the analyses to examine the sensitivity of the results to modifying some of the assumptions of the analysis method. The asymptotic standard errors for the estimates for natural mortality are given in parenthesis. The average proportions by area are given in parenthesis under the estimates for the JARPA qs.


| Case | $b_{\text {rec }}$ |  |  |  | $N_{\text {tot }}$ |  | $\begin{gathered} \hline N_{\text {End-5,5,1 }} / \\ N_{1968,1} \end{gathered}$ | $K_{1930}$ | $\begin{gathered} K_{2960} \\ K_{1930} \\ (\%) \\ \hline(\%) \end{gathered}$ | $\begin{gathered} K_{2000}{ }^{\left(K_{1960}\right.} \\ (\%) \\ \hline(\%) \end{gathered}$ | Natural mortality (ages) |  |  | JARPA $q /$ <br> Mean proportion |  | $\begin{gathered} \hline \text { MSYR } \\ (1+) \end{gathered}$ |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | 1945-68 | 1968-88 | 1988-End | 1945-68 |  |  |  |  |  |  | 0-3 | 10-30 | 35+ | V-E | VI-W |  |
| Reference | 3.608 | -1.696 | 0.322 | 4.338 | 0.518 | -0.499 | 0.671 | 36215 | 1000 | 45 | $\begin{gathered} \hline 0.089 \\ (0.009) \\ \hline \end{gathered}$ | $\begin{gathered} \hline 0.039 \\ (0.003) \\ \hline \end{gathered}$ | $\begin{gathered} 0.106 \\ (0.003) \end{gathered}$ | $\begin{gathered} 0.491 \\ (0.773) \end{gathered}$ | $\begin{gathered} \hline 0.405 \\ (0.227) \\ \hline \end{gathered}$ | 0.036 |
| P1** | -1.501 | -6.501 | -2.101 | -0.985 | -5.188 | -2.981 | 0.198 | 762400 | 8.0 | 0.0 | 0.487 | 0.075 | 0.159 | $\begin{gathered} \hline 4.424 \\ (0.817) \end{gathered}$ | $\begin{aligned} & \hline 5.430 \\ & (0.183 \end{aligned}$ | 0.000 |
| P2 | 3.617 | -1.667 | 0.329 | 4.355 | 0.540 | -0.483 | 0.679 | 35976 | 1000 | 45.7 | $\begin{gathered} 0.088 \\ (0.008) \end{gathered}$ | $\begin{gathered} 0.038 \\ (0.003) \end{gathered}$ | $\begin{aligned} & 0.106 \\ & (0.002) \end{aligned}$ | $\begin{gathered} 0.455 \\ (0.79) \end{gathered}$ | $\begin{gathered} 0.412 \\ (0.211) \end{gathered}$ | 0.036 |
| P3 |  |  |  |  |  |  |  |  |  |  | 0.193 | 0.022 | 0.083 | 0.547 | 0.386 | 0.046 |
| P4 | 4.326 | 0.282 | 0.212 | 4.796 | 2.079 | 0.389 | 0.977 | 19606 | 1000 | 100 | (0.009) | (0.002) | (0.002) | (0.752) | (0.248) | 0.028 |
|  | 3.248 | -2.519 | 0.058 | 3.792 | -0.031 | -0.984 | 0.548 | 52054 | 1000 | 28.0 | (0.007) | (0.002) | (0.002) | (0.784) | (0.216) |  |
| P5 |  |  |  |  |  |  |  |  |  |  | 0.094 | 0.035 | 0.131 | 0.495 | 0.404 | 0.002 |
| P6 | 1.409 | -1.001 | 0.286 | 0.461 | 0.013 | -0.455 | 0.747 | 166325 | 100 | 100 | 0.007 0.116 | ${ }_{0}^{(0.002)}$ | $(0.002)$ 0.125 0 | ${ }_{0}^{(0.772)}$ | $(0.228)$ 0.502 | 0.215 |
|  | 3.254 | -2.248 | -0.884 | 2.756 | -0.167 | -1.598 | 0.419 | 116732 | 159.4 | 90.6 | (0.012) | (0.005) | (0.004) | (0.795) | (0.205) |  |
| P7 | 3.497 | -2.08 | -0.425 | 4.492 | 0.237 | -1.146 | 0.523 | 35916 | 1000 | 38.5 | $\begin{gathered} 0.099 \\ (0.009) \end{gathered}$ | $\begin{gathered} 0.041 \\ (0.004) \end{gathered}$ | $\begin{gathered} 0.115 \\ (0.004) \end{gathered}$ | $\begin{gathered} 0.498 \\ (0.789) \end{gathered}$ | $\begin{gathered} 0.460 \\ (0.211) \end{gathered}$ | 0.040 |
|  | 3.607 | -1.699 | 0.323 | 4.337 | 0.516 | -0.5 | 0.671 | 36256 | 1000 | 44.9 | (0.009) | (0.003) | (0.003) | (0.773) | (0.227) |  |
| S1 | 2.928 | -0.555 | 0.452 | 3.315 | 0.703 | 0.306 | 0.965 | 44074.7 | 41.5 | 1000.0 | $\begin{gathered} 0.198 \\ (0.018) \end{gathered}$ | $\begin{gathered} 0.049 \\ (0.005) \end{gathered}$ | $\begin{aligned} & 0.092 \\ & (0.004) \end{aligned}$ | $\begin{gathered} 0.800 \\ (0.750) \end{gathered}$ | $\begin{gathered} 0.561 \\ (0.200) \end{gathered}$ | 0.025 |
| S2 | 3.603 | -1.738 | 0.3 | 4.32 | 0.483 | -0.523 | 0.663 | 36782 | 1000 | 44.1 | $\begin{gathered} 0.091 \\ (0.009) \end{gathered}$ | $\begin{gathered} 0.039 \\ (0.003) \end{gathered}$ | $\begin{gathered} 0.105 \\ (0.003) \end{gathered}$ | $\begin{gathered} 0.492 \\ (0.774) \end{gathered}$ | $\begin{gathered} 0.407 \\ (0.226) \end{gathered}$ | 0.036 |
| S3 |  |  |  |  |  |  |  |  |  |  | 0.079 | 0.038 | 0.106 | 0.457 | 0.376 | 0.036 |
| S4 | 3.635 | -1.627 | 0.323 | 4.373 | 0.597 | -0.486 | 0.675 | 35409 | 46.3 | 1000 | $(0.008)$ 0.084 | $(0.003)$ 0.035 | $(0.006)$ 0.131 | $(0.773)$ 0.461 | $(0.227)$ 0.377 | 0.002 |
|  | 1.412 | -0.987 | 0.278 | 0.452 | 0.033 | -0.470 | 0.745 | 166262 | 100 | 100 | (0.007) | (0.002) | (0.004) | (0.772) | (0.228) |  |
| S5 |  |  |  |  |  |  |  |  |  |  | 0.091 | 0.039 | ${ }^{0.106}$ | 0.300 | 0.247 | 0.036 |
| S6 | 3.607 | -1.673 | 0.323 | 4.356 | 0.536 | -0.489 | 0.674 | 59130 | 1000 | 45.5 | $(0.009)$ 0.090 | 0.039 | $0.003)$ 0.106 | 0.397 0.3 | 0.327 0.327 | 0.036 |
|  | 3.608 | -1.683 | 0.323 | 4.347 | 0.528 | -0.493 | 0.673 | 44790 | 1000 | 45.3 | (0.009) | (0.003) | (0.003) | (0.773) | (0.227) |  |
| S7 |  |  |  |  |  |  |  |  |  |  | 0.089 | 0.039 | 0.107 |  |  | 0.035 |
|  | 3.581 | -1.781 | 0.280 | 4.289 | 0.461 | $-0.550$ | 0.643 | 37223 | 43.0 | 1000 | (0.009) | (0.003) | (0.006) | (0.807) | (0.193) |  |
| S8 | 3.500 | -2.318 | 0.129 | 4.049 | 0.296 | -0.659 | 0.622 | 43280 | 1000 | 36.2 | $\begin{gathered} 0.105 \\ (0.008) \end{gathered}$ | $\begin{gathered} 0.039 \\ (0.003) \end{gathered}$ | $\begin{gathered} 0.108 \\ (0.003) \end{gathered}$ | $\begin{gathered} 0.507 \\ (0.777) \end{gathered}$ | $\begin{gathered} 0.431 \\ (0.223) \end{gathered}$ | 0.032 |
| S9 |  |  |  |  |  |  |  |  |  |  | 0.087 | 0.040 | 0.110 |  |  | 0.032 |
|  | 3.465 | -2.136 | 0.232 | 4.088 | 0.225 | $-0.748$ | 0.602 | 24575 | 35.9 | 1000 | (0.009) | (0.003) | (0.006) | (0.805) | (0.195) |  |
| S10 | 3.533 | -1.927 | 0.458 | 4.271 | 0.323 | -0.558 | 0.658 | 38453.9 | 41.6 | 1000.0 | $\begin{gathered} 0.099 \\ (0.009) \\ \hline \end{gathered}$ | $\begin{gathered} 0.040 \\ (0.002) \\ \hline \end{gathered}$ | $\begin{gathered} 0.108 \\ (0.006) \\ \hline \end{gathered}$ | $\begin{gathered} 0.512 \\ (0.783) \\ \hline \end{gathered}$ | $\begin{gathered} 0.451 \\ (0.217) \\ \hline \end{gathered}$ | 0.035 |
| D1 | 3.402 | -2.797 | 0.631 | 3.964 | -0.067 | -0.778 | 0.599 | 47577 | 1000 | 31.7 | $\begin{gathered} \hline 0.128 \\ (0.011) \end{gathered}$ | $\begin{gathered} \hline 0.040 \\ (0.003) \end{gathered}$ | $\begin{gathered} 0.111 \\ (0.003) \end{gathered}$ | $\begin{gathered} \hline 0.490 \\ (0.811) \end{gathered}$ | $\begin{gathered} 0.409 \\ (0.189) \end{gathered}$ | 0.031 |
| D2 |  |  |  |  |  |  |  |  |  |  | 0.360 | 0.074 | 0.146 | 3.277 | 3.167 | 0.028 |
|  | 3.553 | -4.120 | -0.356 | 3.662 | -2.480 | -1.055 | 0.516 | 78510 | 1000 | 17.0 | (0.025) | (0.006) | (0.003) | (0.789) | (0.211) |  |
| D3 | 3.602 | -1.710 | 0.319 | 4.337 | 0.507 | -0.505 | 0.669 | 36327 | 1000 | 44.8 | $\begin{gathered} 0.088 \\ (0.009) \end{gathered}$ | $\begin{gathered} 0.039 \\ (0.003) \end{gathered}$ | $\begin{gathered} 0.106 \\ (0.003) \end{gathered}$ | $\begin{gathered} 0.491 \\ (0.773) \end{gathered}$ | $\begin{aligned} & 0.403 \\ & (0.227) \end{aligned}$ | 0.036 |
| D4 |  |  |  |  |  |  |  |  |  |  | 0.089 | 0.039 | 0.106 | 0.491 | 0.405 | 0.036 |
|  | 3.607 | -1.698 | 0.322 | 4.337 | 0.518 | -0.499 | 0.671 | 36234 | 1000 | 45.0 | (0.009) | (0.003) | (0.003) | (0.773) | (0.227) |  |
| D5 | 3.462 | -2.024 | 0.195 | 4.222 | 0.367 | -0.425 | 0.692 | 38143 | 1000 | 43.8 | $\begin{gathered} 0.079 \\ (0.008) \end{gathered}$ | 0.040 $(0.002)$ | 0.105 $(0.002)$ | $\begin{gathered} 0.478 \\ (0.773) \end{gathered}$ | $\begin{gathered} 0.394 \\ (0.227) \end{gathered}$ | 0.035 |
| D6 |  |  |  |  |  |  |  |  |  |  | 0.095 | 0.038 | 0.107 | 0.498 | 0.414 | 0.037 |
|  | 3.785 | -1.591 | 0.250 | 4.498 | 0.602 | -0.577 | 0.652 | 34556 | 1000 | 46.1 | (0.010) | (0.003) | (0.003) | (0.774) | (0.226) |  |
| D7 |  |  |  |  |  |  |  |  |  |  | 0.088 | 0.039 | 0.106 | 0.456 | 0.451 | 0.036 |
|  | 3.604 | -1.706 | 0.317 | 4.332 | 0.513 | -0.506 | 0.671 | 36280 | 1000 | 44.7 | (0.009) | (0.003) | (0.003) | (0.801) | (0.199) |  |
| D8 |  |  |  |  |  |  |  |  |  |  | ${ }_{0}^{0.089}$ | ${ }_{0}^{0.038}$ | ${ }_{0}^{0.106}$ | ${ }_{0}^{0.524}$ | 0.348 | 0.036 |
| D9 | 3.620 | -1.659 | 0.334 | 4.360 | 0.542 | -0.476 | 0.676 | 36573 | 1000 | 45.9 | $(0.009)$ 0.089 | $(0.003)$ 0.039 | $(0.003)$ 0.106 | $(0.743)$ 0.491 | $(0.257)$ 0.405 | 0.036 |
|  | 3.608 | -1.696 | 0.322 | 4.338 | 0.518 | -0.499 | 0.671 | 36215 | 1000 | 45.0 | (0.009) | (0.003) | (0.003) | (0.773) | (0.227) |  |
| D10 |  |  |  |  |  |  |  |  |  |  | 0.089 | 0.039 | 0.106 | 0.491 | 0.405 | 0.036 |
|  | 3.608 | -1.696 | 0.322 | 4.338 | 0.518 | -0.499 | 0.671 | 36215 | 1000 | 45.0 | (0.009) | (0.003) | (0.003) | (0.773) | (0.227) |  |
| D11 |  |  |  |  |  |  |  |  |  |  | 0.086 | 0.042 | 0.111 | 0.499 | 0.424 | 0.035 |
|  | 3.585 | -2.158 | 0.284 | 4.323 | 0.194 | -0.721 | 0.590 | 39795 | 1000 | 39.3 | (0.007) | (0.002) | (0.002) | (0.777) | (0.223) |  |
| D12 | 3.611 | -1.311 | 0.360 | 4.410 | 0.752 | -0.301 | 0.751 | 32665 | 1000 | 51.6 | 0.091 | 0.037 | $0.102$ | $0.493$ <br> (0.769) | $\begin{gathered} 0.394 \\ (0.231) \end{gathered}$ | 0.037 |
| D13 |  |  |  |  |  |  |  |  |  |  | 0.064 | 0.043 | 0.106 | 0.488 | 0.401 | 0.038 |
|  | 3.505 | -1.884 | 0.411 | 4.496 | 0.335 | -0.537 | 0.626 | 35775 | 1000 | 45.0 | (0.007) | (0.002) | (0.002) | (0.772) | (0.228) |  |
| D14 | 3.655 | -1.626 | 0.268 | 4.235 | 0.629 | -0.469 | 0.700 | 36542 | 1000 | 44.9 | $\begin{gathered} 0.105 \\ (0.009) \end{gathered}$ | $\begin{gathered} 0.035 \\ (0.003) \end{gathered}$ | $\begin{gathered} 0.106 \\ (0.003) \end{gathered}$ | $\begin{gathered} 0.487 \\ (0.773) \end{gathered}$ | $\begin{aligned} & 0.401 \\ & (0.277) \end{aligned}$ | 0.035 |

## Table 9

Negative log-likelihood for the model configurations which are comparable with the reference case model (the values in parenthesis are numbers of estimated parameters). Values indicated by asterisks are for model configurations for which the Hessian matrix was not positive definite. Results are not shown for model configurations S2 and S7 because the weightings of the likelihood function for these two configurations are not compatible with that for the reference case model.

| Model Configuration | Negative log-likelihood | Model Configuration | Negative log-likelihood |
| :---: | :---: | :---: | :---: |
| Reference | $20499.6(1039)$ | S1 | $21200.6(1039)$ |
| P1 | $20673.6(1039)^{*}$ | S3 | $20500.2(1035)$ |
| P2 | $20532.3(944)$ | S4 | $20626.4(1031)$ |
| P3 | $22266.2(749)$ | S5 | $20497.7(1039)$ |
| P4 | $20503.7(1038)$ | S6 | $20498.5(1039)$ |
| P5 | $20626.0(1035)$ | S8 | $21333.0(947)$ |
| P6 | $20446.9(1199)$ | S9 | $20532.2(1034)$ |
| P7 | $20494.8(1041)$ | $20496.1(1048)$ |  |
| P8 | $20499.5(1039)$ |  |  |



Figure 1. Time-trajectories of total (1+) population size, recruitment (from 1930 and from 1975), and recruitment deviations. The dotted lines indicate $95 \%$ asymptotic confidence intervals.


Figure 2. Age-specific natural mortality by stock, and time-trajectories of carrying capacity, the number of mature females, and total (1+) population size relative to carrying capacity. The dotted lines indicate 95\% asymptotic confidence intervals.


Figure 3. Time-trajectories of the proportion of the total population in each of the five areas considered in the model as well as those of the growth rate parameter $k$ (by stock and sex).


Figure 4. Time-trajectories of the number of calves-per-mature female by stock.


Figure 5. Fits to the indices of relative abundance (from JARPA). The bars indicate $95 \%$ confidence intervals based on the supplied sampling standard errors.


Figure 6. Fits to the indices of absolute abundance (from IDCR). The bars indicate 95\% confidence intervals based on the supplied sampling standard errors.



Figure 7. Summaries of the fits to the length-frequency data for each fleet and sex. The left panels show the standardized residuals by size-class and year in the form of a bubble plot, the centre panel plots the observed sample sizes (number of animals actually measured) against the computed effective sample size, and the right panel shows the fit of the model to the sum over years of the number of measured animals. The solid lines in the center panels are 1-1 lines, the dashed lines are the best linear relationships between observed and modelinferred expected sample sizes, and the dotted lines are the assumed relationships between observed and modeliinferred expected sample sizes (a slope of 0.7).


Figure 8. Detailed results of the fits to the length-frequency data for one fleet and sex. In the histograms, the bars indicate the raw data and the dots the model fits. The last three panels are the overall summaries of the fits.


Figure 9. Observed (dots) and model-predicted (lines) mean ages versus length (ft) for one fleet-sex combination.

Stock $=1 ;$ Sex $=F$
Stock = I; Sex = M


Figure 10. Length-at-age versus time by sex and stock.


Figure 11. Length-selectivity patterns for five of the ten fleets (the Japanese fleets). Results are not shown for the ex-USSR fleets are they are identical to those for the Japanese fleets.


Figure 12. Age-selectivity patterns for five of the ten fleets (the Japanese fleets). Results are not shown for the ex-USSR fleets are they are identical to those for the Japanese fleets.


Figure 13. The time-trajectory of $1+$ population size for the reference case analysis and the " P " sensitivity tests (" 0 " denotes the reference case analysis). Results are shown for the I stock in the left panels and for the P stock in the right panels.


Figure 14. The time-trajectory of $1+$ population size for the reference case analysis and the " S " sensitivity tests ("0" denotes the reference case analysis). Results are shown for the I stock in the left panels and for the P stock in the right panels. Sensitivity test S7 did lead to a positive definite Hessian matrix.


Figure 15. The time-trajectory of $1+$ population size for the reference case analysis and the " D " sensitivity tests ("0" denotes the reference case analysis). Results are shown for the I stock in the left panels and for the P stock in the right panels.


[^0]:    ${ }^{1}$ Age-reading error has been quantified for Antarctic minke whales by Kitakado and Punt (2010, 2013).

[^1]:    ${ }^{2}$ Two Japanese 'fleets' are considered so that the data for commercial and Scientific Permit catches can be treated separately.

[^2]:    ${ }^{3}$ As calves are not harvested, this formulation for density-dependence conceptually encompasses densitydependent effects in the survival rate of calves.

[^3]:    ${ }^{4}$ Note that very high weight is assigned to this component of the objective function so the model effectively replicates the actual catches exactly.

[^4]:    ${ }^{5}$ Note that the evaluation of the impact of age-reading error is determined before the application of the plus-group.

[^5]:    ${ }^{6}$ Additional variance is implicitly accounted for by allowing the proportion of the population in each area to change over time.

[^6]:    ${ }^{7}$ AIC should be interpreted with some caution given the somewhat ad hoc way the various data sources are weighted.

[^7]:    ${ }^{8}$ The analyses started from a parameter vector which was very far from the true minimum and changes to those values might resolve these non-convergence problems.

